

# CS70: Jean Walrand: Lecture 34.

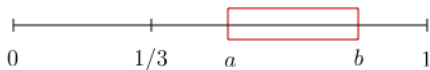
## Continuous Probability 1

1. Examples
2. Events
3. Continuous Random Variables

## Uniformly at Random in $[0, 1]$ .

Choose a real number  $X$ , uniformly at random in  $[0, 1]$ .

What is the probability that  $X$  is exactly equal to  $1/3$ ? Well, ..., 0.



What is the probability that  $X$  is exactly equal to 0.6? Again, 0.

In fact, for any  $x \in [0, 1]$ , one has  $Pr[X = x] = 0$ .

How should we then describe ‘choosing uniformly at random in  $[0, 1]$ ’?

Here is the way to do it:

$$Pr[X \in [a, b]] = b - a, \forall 0 \leq a \leq b \leq 1.$$

Makes sense:  $b - a$  is the fraction of  $[0, 1]$  that  $[a, b]$  covers.

## Uniformly at Random in $[0, 1]$ .

Let  $[a, b]$  denote the **event** that the point  $X$  is in the interval  $[a, b]$ .

$$\Pr[[a, b]] = \frac{\text{length of } [a, b]}{\text{length of } [0, 1]} = \frac{b - a}{1} = b - a.$$

Intervals like  $[a, b] \subseteq \Omega = [0, 1]$  are **events**.

More generally, events in this space are **unions of intervals**.

Example: the event  $A$  - “within 0.2 of 0 or 1” is  $A = [0, 0.2] \cup [0.8, 1]$ .

Thus,

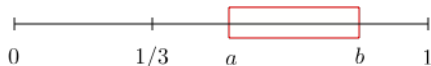
$$\Pr[A] = \Pr[[0, 0.2]] + \Pr[[0.8, 1]] = 0.4.$$

More generally, if  $A_n$  are pairwise disjoint intervals in  $[0, 1]$ , then

$$\Pr[\cup_n A_n] := \sum_n \Pr[A_n].$$

Many subsets of  $[0, 1]$  are of this form. Thus, the probability of those sets is well defined. We call such sets **events**.

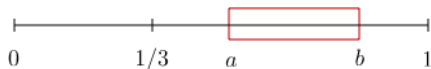
## Uniformly at Random in $[0, 1]$ .



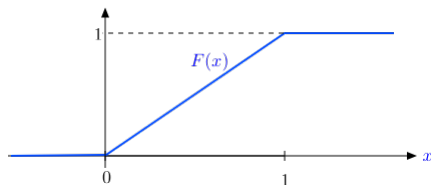
Note: A **radical** change in approach. For a finite probability space,  $\Omega = \{1, 2, \dots, N\}$ , we started with  $Pr[\omega] = p_\omega$ . We then defined  $Pr[A] = \sum_{\omega \in A} p_\omega$  for  $A \subset \Omega$ . We used the same approach for countable  $\Omega$ .

For a continuous space, e.g.,  $\Omega = [0, 1]$ , we cannot start with  $Pr[\omega]$ , because this will typically be 0. Instead, we start with  $Pr[A]$  for some events  $A$ . Here, we started with  $A =$  interval, or union of intervals.

## Uniformly at Random in $[0, 1]$ .

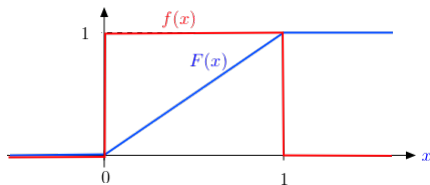


Note:  $Pr[X \leq x] = x$  for  $x \in [0, 1]$ . Also,  $Pr[X \leq x] = 0$  for  $x < 0$  and  $Pr[X \leq x] = 1$  for  $x > 1$ . Let us define  $F(x) = Pr[X \leq x]$ .



Then we have  $Pr[X \in (a, b]] = Pr[X \leq b] - Pr[X \leq a] = F(b) - F(a)$ .  
Thus,  $F(\cdot)$  specifies the probability of all the events!

## Uniformly at Random in $[0, 1]$ .



$$\Pr[X \in (a, b]] = \Pr[X \leq b] - \Pr[X \leq a] = F(b) - F(a).$$

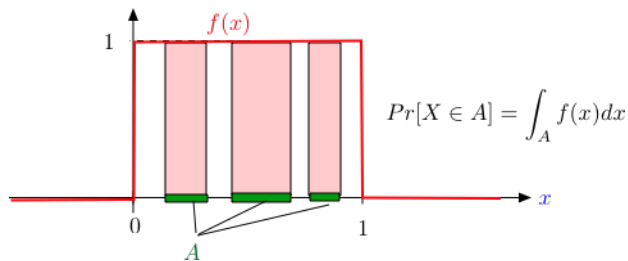
An alternative view is to define  $f(x) = \frac{d}{dx} F(x) = 1\{x \in [0, 1]\}$ . Then

$$F(b) - F(a) = \int_a^b f(x) dx.$$

Thus, the probability of an event is the integral of  $f(x)$  over the event:

$$\Pr[X \in A] = \int_A f(x) dx.$$

## Uniformly at Random in $[0, 1]$ .



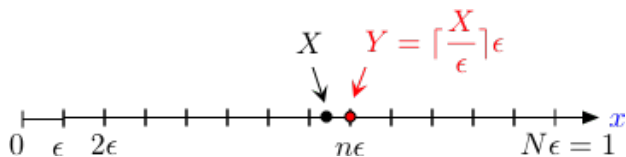
Think of  $f(x)$  as describing how  
one unit of probability is spread over  $[0, 1]$ : uniformly!

Then  $Pr[X \in A]$  is the probability mass over  $A$ .

Observe:

- ▶ This makes the probability automatically additive.
- ▶ We need  $f(x) \geq 0$  and  $\int_{-\infty}^{\infty} f(x) dx = 1$ .

## Uniformly at Random in $[0, 1]$ .



**Discrete Approximation:** Fix  $N \gg 1$  and let  $\epsilon = 1/N$ .

Define  $Y = n\epsilon$  if  $(n-1)\epsilon < X \leq n\epsilon$  for  $n = 1, \dots, N$ .

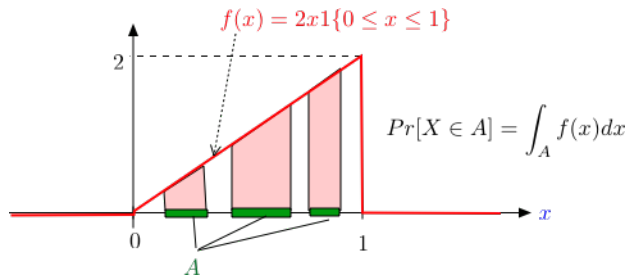
Then  $|X - Y| \leq \epsilon$  and  $Y$  is discrete:  $Y \in \{\epsilon, 2\epsilon, \dots, N\epsilon\}$ .

Also,  $\Pr[Y = n\epsilon] = \frac{1}{N}$  for  $n = 1, \dots, N$ .

Thus,  $X$  is 'almost discrete.'



## Nonuniformly at Random in $[0, 1]$ .



This figure shows a different choice of  $f(x) \geq 0$  with  $\int_{-\infty}^{\infty} f(x) dx = 1$ .

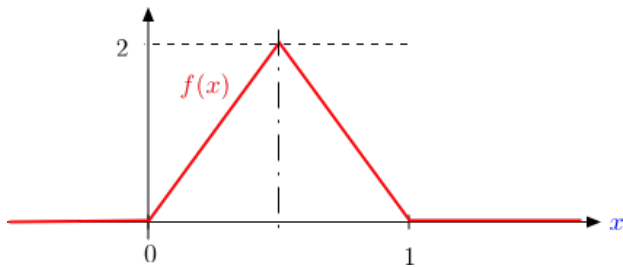
It defines another way of choosing  $X$  at random in  $[0, 1]$ .

Note that  $X$  is more likely to be closer to 1 than to 0.

One has  $Pr[X \leq x] = \int_{-\infty}^x f(u) du = x^2$  for  $x \in [0, 1]$ .

Also,  $Pr[X \in (x, x + \varepsilon)] = \int_x^{x+\varepsilon} f(u) du \approx f(x)\varepsilon$ .

## Another Nonuniform Choice at Random in $[0, 1]$ .



This figure shows yet a different choice of  $f(x) \geq 0$  with  $\int_{-\infty}^{\infty} f(x) dx = 1$ .

It defines another way of choosing  $X$  at random in  $[0, 1]$ .

Note that  $X$  is more likely to be closer to  $1/2$  than to  $0$  or  $1$ .

For instance,  $Pr[X \in [0, 1/3]] = \int_0^{1/3} 4x dx = 2[x^2]_0^{1/3} = \frac{2}{9}$ .

Thus,  $Pr[X \in [0, 1/3]] = Pr[X \in [2/3, 1]] = \frac{2}{9}$  and  $Pr[X \in [1/3, 2/3]] = \frac{5}{9}$ .

## General Random Choice in $\mathfrak{R}$

Let  $F(x)$  be a nondecreasing function with  $F(-\infty) = 0$  and  $F(+\infty) = 1$ .

Define  $X$  by  $Pr[X \in (a, b]] = F(b) - F(a)$  for  $a < b$ . Also, for  $a_1 < b_1 < a_2 < b_2 < \dots < b_n$ ,

$$\begin{aligned} Pr[X \in (a_1, b_1] \cup (a_2, b_2] \cup (a_n, b_n)] \\ &= Pr[X \in (a_1, b_1]] + \dots + Pr[X \in (a_n, b_n)] \\ &= F(b_1) - F(a_1) + \dots + F(b_n) - F(a_n). \end{aligned}$$

Let  $f(x) = \frac{d}{dx} F(x)$ . Then,

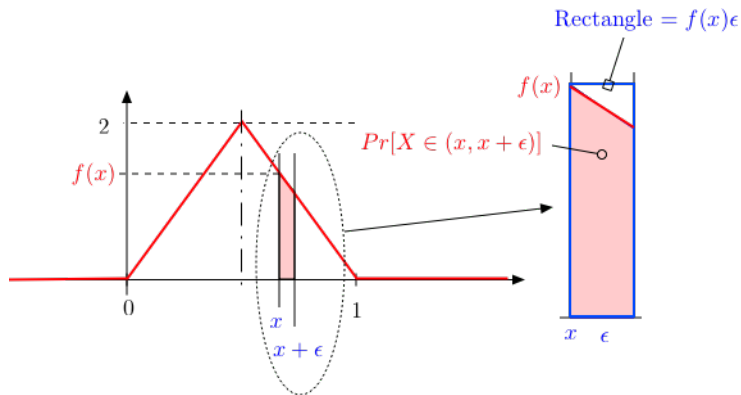
$$Pr[X \in (x, x + \varepsilon]] = F(x + \varepsilon) - F(x) \approx f(x)\varepsilon.$$

Here,  $F(x)$  is called the **cumulative distribution function (cdf)** of  $X$  and  $f(x)$  is the **probability density function (pdf)** of  $X$ .

To indicate that  $F$  and  $f$  correspond to the RV  $X$ , we will write them  $F_X(x)$  and  $f_X(x)$ .

$$Pr[X \in (x, x + \epsilon)]$$

An illustration of  $Pr[X \in (x, x + \epsilon)] \approx f_X(x)\epsilon$ :



Thus, the pdf is the 'local probability by unit length.'

It is the 'probability density.'

# Discrete Approximation

Fix  $\varepsilon \ll 1$  and let  $Y = n\varepsilon$  if  $X \in (n\varepsilon, (n+1)\varepsilon]$ .

Thus,  $Pr[Y = n\varepsilon] = F_X((n+1)\varepsilon) - F_X(n\varepsilon)$ .

Note that  $|X - Y| \leq \varepsilon$  and  $Y$  is a discrete random variable.

Also, if  $f_X(x) = \frac{d}{dx} F_X(x)$ , then  $F_X(x + \varepsilon) - F_X(x) \approx f_X(x)\varepsilon$ .

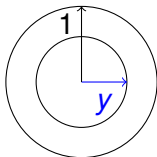
Hence,  $Pr[Y = n\varepsilon] \approx f_X(n\varepsilon)\varepsilon$ .

Thus, we can think of  $X$  of being almost discrete with

$Pr[X = n\varepsilon] \approx f_X(n\varepsilon)\varepsilon$ .

## Example: CDF

Example: hitting random location on gas tank.  
Random location on circle.



Random Variable:  $Y$  distance from center.  
Probability within  $y$  of center:

$$\begin{aligned}Pr[Y \leq y] &= \frac{\text{area of small circle}}{\text{area of dartboard}} \\ &= \frac{\pi y^2}{\pi} = y^2.\end{aligned}$$

Hence,

$$F_Y(y) = Pr[Y \leq y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \leq y \leq 1 \\ 1 & \text{for } y > 1 \end{cases}$$

## Calculation of event with dartboard..

Probability between .5 and .6 of center?

Recall CDF.

$$F_Y(y) = Pr[Y \leq y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \leq y \leq 1 \\ 1 & \text{for } y > 1 \end{cases}$$

$$\begin{aligned} Pr[0.5 < Y \leq 0.6] &= Pr[Y \leq 0.6] - Pr[Y \leq 0.5] \\ &= F_Y(0.6) - F_Y(0.5) \\ &= .36 - .25 \\ &= .11 \end{aligned}$$

# PDF.

Example: “Dart” board.

Recall that

$$F_Y(y) = Pr[Y \leq y] = \begin{cases} 0 & \text{for } y < 0 \\ y^2 & \text{for } 0 \leq y \leq 1 \\ 1 & \text{for } y > 1 \end{cases}$$

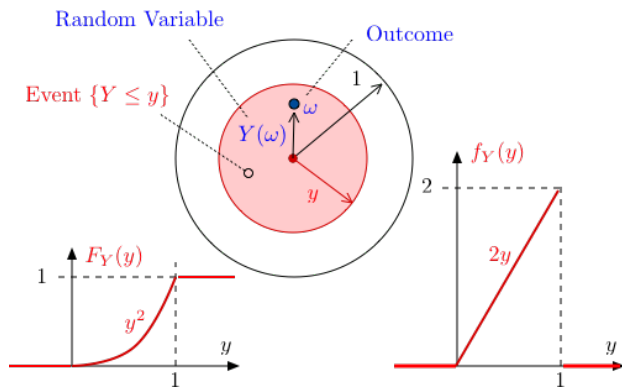
$$f_Y(y) = F'_Y(y) = \begin{cases} 0 & \text{for } y < 0 \\ 2y & \text{for } 0 \leq y \leq 1 \\ 0 & \text{for } y > 1 \end{cases}$$

The cumulative distribution function (cdf) and probability distribution function (pdf) give full information.

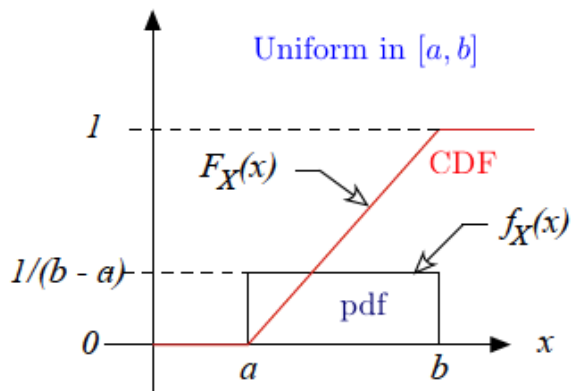
Use whichever is convenient.



# Target



$U[a, b]$

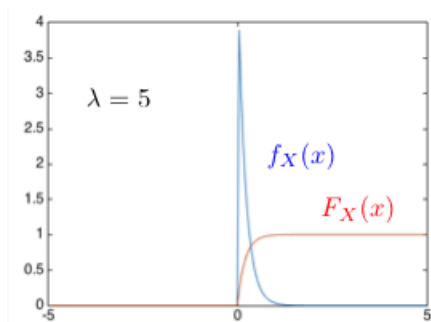
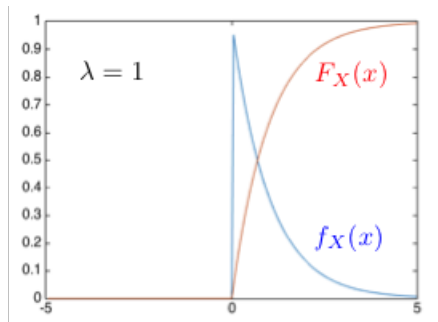


## Expo( $\lambda$ )

The exponential distribution with parameter  $\lambda > 0$  is defined by

$$f_X(x) = \lambda e^{-\lambda x} 1_{\{x \geq 0\}}$$

$$F_X(x) = \begin{cases} 0, & \text{if } x < 0 \\ 1 - e^{-\lambda x}, & \text{if } x \geq 0. \end{cases}$$



Note that  $Pr[X > t] = e^{-\lambda t}$  for  $t > 0$ .

# Random Variables

Continuous random variable  $X$ , specified by

1.  $F_X(x) = Pr[X \leq x]$  for all  $x$ .

**Cumulative Distribution Function (cdf).**

$$Pr[a < X \leq b] = F_X(b) - F_X(a)$$

1.1  $0 \leq F_X(x) \leq 1$  for all  $x \in \mathfrak{R}$ .

1.2  $F_X(x) \leq F_X(y)$  if  $x \leq y$ .

2. Or  $f_X(x)$ , where  $F_X(x) = \int_{-\infty}^x f_X(u) du$  or  $f_X(x) = \frac{d(F_X(x))}{dx}$ .

**Probability Density Function (pdf).**

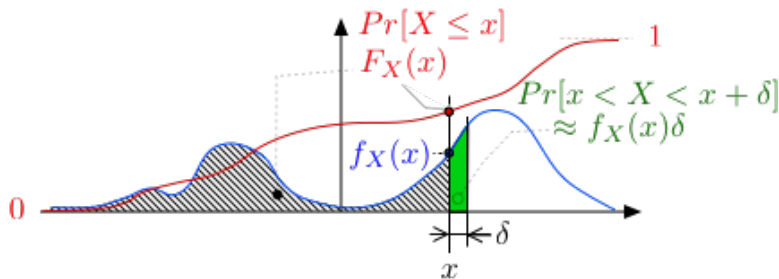
$$Pr[a < X \leq b] = \int_a^b f_X(x) dx = F_X(b) - F_X(a)$$

2.1  $f_X(x) \geq 0$  for all  $x \in \mathfrak{R}$ .

2.2  $\int_{-\infty}^{\infty} f_X(x) dx = 1$ .

Recall that  $Pr[X \in (x, x + \delta)] \approx f_X(x)\delta$ . Think of  $X$  taking discrete values  $n\delta$  for  $n = \dots, -2, -1, 0, 1, 2, \dots$  with  $Pr[X = n\delta] = f_X(n\delta)\delta$ .

## A Picture



The pdf  $f_X(x)$  is a nonnegative function that integrates to 1.  
The cdf  $F_X(x)$  is the integral of  $f_X$ .

$$Pr[x < X < x + \delta] \approx f_X(x)\delta$$

$$Pr[X \leq x] = F_X(x) = \int_{-\infty}^x f_X(u)du$$

# Summary

## Continuous Probability 1

1. pdf:  $Pr[X \in (x, x + \delta]] = f_X(x)\delta$ .
2. CDF:  $Pr[X \leq x] = F_X(x) = \int_{-\infty}^x f_X(y)dy$ .
3.  $U[a, b]$ :  $f_X(x) = \frac{1}{b-a}1\{a \leq x \leq b\}$ ;  $F_X(x) = \frac{x-a}{b-a}$  for  $a \leq x \leq b$ .
4.  $Expo(\lambda)$ :  
 $f_X(x) = \lambda \exp\{-\lambda x\}1\{x \geq 0\}$ ;  $F_X(x) = 1 - \exp\{-\lambda x\}$  for  $x \geq 0$ .
5. Target:  $f_X(x) = 2x1\{0 \leq x \leq 1\}$ ;  $F_X(x) = x^2$  for  $0 \leq x \leq 1$ .